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Optimal imputation methods under stratified ranked set sampling

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Abstract:

- It is long familiar that the stratified ranked set sampling ($SRSS$) is more efficient than ranked set sampling (RSS) and stratified random sampling ($StRS$). The existence of missing values may alter the final inference of any study. This paper is a fundamental effort to suggest some combined and separate imputation methods in the presence of missing data under $SRSS$. The proposed imputation methods become superior than the mean imputation method, ratio imputation method, Diana and Perri ([13]) type imputation method, and Sohail et al. ([32]) type imputation methods. A simulation study is administered over two hypothetically drawn asymmetric populations.

Keywords:

- *Missing values; imputation; stratified ranked set sampling.*

AMS Subject Classification:

- 62D05.

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1. INTRODUCTION

The dilemma of missing value is very usual in a sample survey and its presence can spoil the traditional results. Therefore, it becomes essential to resolve the problem of missing values in a data set. The well-known imputation technique is used to replace the missing values. Three basic concepts on missing values were suggested by Rubin ([23]), such as missing at random (MAR), observed at random (OAR), and parameter distribution (PD). Several renowned authors like Lee *et al.* ([17]), Singh and Horn ([30]), Singh and Deo ([27]), Singh ([28]), and Singh and Valdes ([31]) introduced various imputation methods in the presence of missing values. Heitzan and Basu ([14]) exhibited a difference between missing at random and missing completely at random ($MCAR$) approach. Thereafter, Ahmed *et al.* ([1]), Kadilar and Cingi ([15]), Diana and Perri ([13]) and Bhushan and Pandey ([4, 5]), Mohamed *et al.* ([21]), Prasad ([22]), Bouza *et al.* ([9]), Bouza-Herrera and Viada ([10]), and Bhushan *et al.* ([6, 7]) utilized $MCAR$ strategy in their study for the imputation of missing values.
In real life, situations may emerge where it is either difficult to measure the study variable or indeed expensive but can be ranked either visually or by any cost free method. In such circumstances, McIntyre ([20]) proposed the idea of ranked set sampling (RSS), which is superior to simple random sampling but did not furnish any mathematical support. Takahasi and Wakimoto ([33]) extended the idea of McIntyre ([20]) and provided the obligatory mathematical foundation to the theory of RSS. Samawi ([25]) envisaged the idea of $SRSS$ superior to $StRS$. Samawi and Siam ([26]) introduced combined and separate ratio estimators under $SRSS$. Mandowara and Mehta ([19]) considered modified ratio estimators under $SRSS$. Linder *et al.* ([18]) investigated the regression estimator under $SRSS$. Khan and Shabbir ([16]) suggested Hartley-Ross type unbiased estimators under RSS and $SRSS$. Recently, Saini and Kumar ([24]) suggested the ratio estimator using quartile as an auxiliary information under $SRSS$.
In sample surveys, when each group contains very small observations, then each observation becomes essential to draw conclusions. Further use of such kind of data set consisting of missing values may vitiate the final conclusion and decrease the efficiency of the estimator as well. In order to tackle with such kind of problems, Bouza and Al-Omari ([8]) suggested mean imputation and ratio methods for the median estimator under RSS. Al-Omari and Bouza ([2]) introduced ratio estimators of the population mean with missing values under RSS. Sohail *et al.* ([32]) suggested ratio type imputation methods under RSS.
In this paper, we suggest some imputation methods in the presence of missing data under $SRSS$. The rest paper is arranged in subsequent sections. In the next section, we discuss the sampling methodology along with the notations used throughout the manuscript. In Section 3, the combined and separate imputation methods are reviewed. In Section 4, we have suggested combined and separate classes of imputation methods. The theoretical comparisons of combined and separate imputation methods are given in Section 5, whereas Section 6 deals with the simulation study conducted in favour of theoretical findings. Lastly, the

conclusion is given in Section 7.

2. Methodology and notations

The procedure of ranked set sampling consists of drawing m simple random samples of size m from the population. These m units are now ranked within each set with respect to the variable of interest, say x. The first smallest unit is quantified from the first set for the measurement of the auxiliary variable along with the associated study variables. The unit with the second smallest rank is quantified from the second ranked set for the measurement of the auxiliary variables along with the associated study variable and the process is carried on as far as the m^{th} smallest unit is quantified from the last set. The above process is known as a cycle. The repetition of this whole procedure up to k times furnishes $n = mk$ ranked set samples.

The stratified ranked set sampling is a sampling procedure analogous to stratified random sampling, which is based on splitting a population into L mutually exclusive and exhaustive strata and a ranked set sample of $n_h = m_h k$ units are measured within each stratum such that $h = 1, 2, ..., L$. The sampling is accomplished independently across the strata. Thus, $SRSS$ scheme can be supposed to a collection of L separate ranked set samples.

Consider a finite population U comprised of N measurable units with values y_i, $i \in U$. Let a stratified ranked set sample of size $n = m_h k$ be chosen from U to estimate the population mean of the study variable y. Let r be the number of responding elements out of n sampled elements. Let P be the probability that i^{th} respondent associated with a responding class A and $(1 - P)$ be the probability that i^{th} respondent associated with the non-responding class $\bar{A}$. Moreover, note that $s = A \cup \bar{A}$ and let the values $y_i,\ i \in A$ be observable for each characteristic, but for the characteristic $i \in \bar{A}$ the values are missing and require imputation in order to establish the complete frame of data to draw a reasonable inference. The auxiliary variable x will be used to execute the imputation of missing values and let the ranking be performed over the auxiliary variables as well.

The succeeding notations would be used from the beginning to end in the case of combined estimators.

Let $\bar{y}_{r,srss} = \bar{Y}(1 + \epsilon_0)$, $\bar{x}_{r,srss} = \bar{X}(1 + \epsilon_1)$, $\bar{x}_{n,srss} = \bar{X}(1 + \epsilon_2)$ such that $E(\epsilon_0) = E(\epsilon_1) = E(\epsilon_2) = 0$ and

$$E(\epsilon_0^2) = \sum_{h=1}^{L} W_h^2 \left(\frac{C_{y_h}^2}{m_h k P} - \frac{1}{m_h^2 k P} \sum_{i=1}^{m_h} \frac{\tau_{y_h}^2}{\bar{Y}^2} \right) = \sum_{h=1}^{L} W_h^2 \left(\gamma^* C_{y_h}^2 - D_{y_h}^{2^*} \right) = I_0^*$$

$$E(\epsilon_1^2) = \sum_{h=1}^{L} W_h^2 \left(\frac{C_{x_h}^2}{m_h k P} - \frac{1}{m_h^2 k P} \sum_{i=1}^{m_h} \frac{\tau_{x_h}^2}{\bar{X}^2} \right) = \sum_{h=1}^{L} W_h^2 \left(\gamma^* C_{x_h}^2 - D_{x_h}^{2^*} \right) = I_1^*$$

$$E(\epsilon_2^2) = \sum_{h=1}^{L} W_h^2 \left(\frac{C_{x_h}^2}{m_h k} - \frac{1}{m_h^2 k} \sum_{i=1}^{m_h} \frac{\tau_{x_h}^2}{\bar{X}^2} \right) = \sum_{h=1}^{L} W_h^2 \left(\gamma C_{x_h}^2 - D_{x_h}^2 \right) = I_1$$

$$E(\epsilon_0, \epsilon_1) = \sum_{h=1}^{L} W_h^2 \left(\frac{\rho_{x_h y_h} C_{x_h} C_{y_h}}{m_h k P} - \frac{1}{m_h^2 k P} \sum_{i=1}^{m_h} \frac{\tau_{x_h y_h}}{\bar{X}\bar{Y}} \right)$$
$$= \sum_{h=1}^{L} W_h^2 \left(\gamma^* \rho_{x_h y_h} C_{x_h} C_{y_h} - D_{x_h y_h}^* \right) = I_{01}^*$$

$$E(\epsilon_0, \epsilon_2) = \sum_{h=1}^{L} W_h^2 \left(\frac{\rho_{x_h y_h} C_{x_h} C_{y_h}}{m_h k} - \frac{1}{m_h^2 k} \sum_{i=1}^{m_h} \frac{\tau_{x_h y_h}}{\bar{X}\bar{Y}} \right)$$
$$= \sum_{h=1}^{L} W_h^2 \left(\gamma \rho_{x_h y_h} C_{x_h} C_{y_h} - D_{x_h y_h} \right) = I_{01}$$

$$E(\epsilon_1, \epsilon_2) = \sum_{h=1}^{L} W_h^2 \left(\frac{C_{x_h}^2}{m_h k} - \frac{1}{m_h^2 k} \sum_{i=1}^{m_h} \frac{\tau_{x_h}^2}{\bar{X}^2} \right) = \sum_{h=1}^{L} W_h^2 \left(\gamma C_{x_h}^2 - D_{x_h}^2 \right) = I_1$$

where $\gamma^* = 1/m_h k P$, $\gamma = 1/m_h k$, $\tau_{y_h} = (\mu_{y_h} - \bar{Y}_h)$, $\tau_{x_h} = (\mu_{x_h} - \bar{X}_h)$ and $\tau_{x_h y_h} = (\mu_{x_h} - \bar{X}_h)(\mu_{y_h} - \bar{Y}_h)$. Also, $C_{x_h} = S_{x_h}/\bar{X}$ and $C_{y_h} = S_{y_h}/\bar{Y}$ are the coefficients of variation of auxiliary variable x and study variable y, respectively. In the case of separate estimators, the following notations will be used throughout the paper.
Let $\bar{y}_{r,h[rss]} = \bar{Y}_h(1 + e_{0_h})$, $\bar{x}_{r,h(rss)} = \bar{X}_h(1 + e_{1_h})$, $\bar{x}_{n,h(rss)} = \bar{X}_h(1 + e_{2_h})$ such that $E(e_{0_h}) = E(e_{1_h}) = E(e_{2_h}) = 0$ and

$$E(e_{0_h}^2) = \left(\frac{C_{y_h}^2}{m_h k P} - \frac{1}{m_h^2 k P} \sum_{i=1}^{m_h} \frac{\tau_{y_h}^2}{\bar{Y}_h^2} \right) = \left(\gamma^* C_{y_h}^2 - M_{y_h}^{2^*} \right) = J_0^*$$

$$E(e_{1_h}^2) = \left(\frac{C_{x_h}^2}{m_h k P} - \frac{1}{m_h^2 k P} \sum_{i=1}^{m_h} \frac{\tau_{x_h}^2}{\bar{X}_h^2} \right) = \left(\gamma^* C_{x_h}^2 - M_{x_h}^{2^*} \right) = J_1^*$$

$$E(e_{2_h}^2) = \left(\frac{C_{x_h}^2}{m_h k} - \frac{1}{m_h^2 k} \sum_{i=1}^{m_h} \frac{\tau_{x_h}^2}{\bar{X}_h^2} \right) = \left(\gamma C_{x_h}^2 - M_{x_h}^2 \right) = J_1$$

$$E(e_{0_h}, e_{1_h}) = \left(\frac{\rho_{x_h y_h} C_{x_h} C_{y_h}}{m_h k P} - \frac{1}{m_h^2 k P} \sum_{i=1}^{m_h} \frac{\tau_{x_h y_h}}{\bar{X}_h \bar{Y}_h} \right) = \left(\gamma^* \rho_{x_h y_h} C_{x_h} C_{y_h} - M_{x_h y_h}^* \right) = J_{01}^*$$

$$E(e_{0_h}, e_{2_h}) = \left(\frac{\rho_{x_h y_h} C_{x_h} C_{y_h}}{m_h k} - \frac{1}{m_h^2 k} \sum_{i=1}^{m_h} \frac{\tau_{x_h y_h}}{\bar{X}_h \bar{Y}_h} \right) = \left(\gamma \rho_{x_h y_h} C_{x_h} C_{y_h} - M_{x_h y_h} \right) = J_{01}$$

$$E(\epsilon_{1_h}, \epsilon_{2_h}) = \left(\frac{C_{x_h}^2}{m_h k} - \frac{1}{m_h^2 k} \sum_{i=1}^{m_h} \frac{\tau_{x_h}^2}{\bar{X}_h^2} \right) = \left(\gamma C_{x_h}^2 - M_{x_h}^2 \right) = J_1$$

where $\tau_{y_h} = (\mu_{y_h} - \bar{Y}_h)$, $\tau_{x_h} = (\mu_{x_h} - \bar{X}_h)$ and $\tau_{x_h y_h} = (\mu_{x_h} - \bar{X}_h)(\mu_{y_h} - \bar{Y}_h)$, $C_{x_h} = S_{x_h}/\bar{X}_h$ and $C_{y_h} = S_{y_h}/\bar{Y}_h$.

3. Recap of imputation methods

In this section, we consider a concise recap of existing prominent combined and separate imputation methods under $SRSS$.

3.1. Combined imputation methods

The mean method of imputation under $SRSS$ is given by

$$y^c_{.i_m} = \begin{cases} y_i & \text{for } i \in A \\ \bar{y}_{r,srss} & \text{for } i \in \bar{A} \end{cases}$$

The sequent estimator is given by

$$T^c_m = \bar{y}_{r,srss}$$

where $\bar{y}_{r,srss} = \sum_{h=1}^{L} W_h \bar{y}_{h[rss]}$ is the stratified ranked set sample mean of study variable y. Also, $W_h = N_h/N$ is the weight of stratum h and N_h and N are the size of stratum h and total population size, respectively.
The imputation methods are categorized into three situations under the availability of auxiliary informations.
Situation I: When $\bar{X}$ is known and $\bar{x}_{n,srss}$ is utilized.
Situation II: When $\bar{X}$ is known and $\bar{x}_{r,srss}$ is utilized.
Situation III: When $\bar{X}$ is unknown and $\bar{x}_{n,srss}$, $\bar{x}_{r,srss}$ are utilized.
The classical combined ratio type imputation methods are defined under $SRSS$ as
Situation I

$$y^c_{.i_{R_1}} = \begin{cases} y_i & \text{for } i \in A \\ \frac{1}{n-r}\left[n\bar{y}_{r,rss}\left(\frac{\bar{X}}{\bar{x}_{n,srss}}\right) - r\bar{y}_{r,srss}\right] & \text{for } i \in \bar{A} \end{cases}$$

Situation II

$$y^c_{.i_{R_2}} = \begin{cases} y_i & \text{for } i \in A \\ \frac{1}{n-r}\left[n\bar{y}_{r,rss}\left(\frac{\bar{X}}{\bar{x}_{r,srss}}\right) - r\bar{y}_{r,srss}\right] & \text{for } i \in \bar{A} \end{cases}$$

Situation III

$$y^c_{.i_{R_3}} = \begin{cases} y_i & \text{for } i \in A \\ \frac{1}{n-r}\left[n\bar{y}_{r,rss}\left(\frac{\bar{x}_{n,srss}}{\bar{x}_{r,srss}}\right) - r\bar{y}_{r,srss}\right] & \text{for } i \in \bar{A} \end{cases}$$

The sequent estimators are

$$T^c_{R_1} = \bar{y}_{r,srss}\left(\frac{\bar{X}}{\bar{x}_{n,srss}}\right)$$
$$T^c_{R_2} = \bar{y}_{r,srss}\left(\frac{\bar{X}}{\bar{x}_{r,srss}}\right)$$
$$T^c_{R_3} = \bar{y}_{r,srss}\left(\frac{\bar{x}_{n,srss}}{\bar{x}_{r,srss}}\right)$$

where $\bar{x}_{n,srss} = \sum_{h=1}^{L} W_h \bar{x}_{h(rss)}$ is the stratified ranked set sample mean of auxiliary variable x.
Following Diana and Perri ([13]), we define the regression imputation methods to impute the missing value under $SRSS$ as
Situation I

$$y^c_{.i_{DP_1}} = \begin{cases} y_i & \text{for } i \in A \\ \bar{y}_r + \frac{n}{n-r} b_1(\bar{X} - \bar{x}_{n,srss}) y_{r,srss} & \text{for } i \in \bar{A} \end{cases}$$

Situation II

$$y^c_{.i_{DP_2}} = \begin{cases} y_i & \text{for } i \in A \\ \bar{y}_r + \frac{n}{n-r} b_2(\bar{X} - \bar{x}_{r,srss}) y_{r,srss} & \text{for } i \in \bar{A} \end{cases}$$

Situation III

$$y^c_{.i_{DP_3}} = \begin{cases} y_i & \text{for } i \in A \\ \bar{y}_r + \frac{n}{n-r} b_3(\bar{x}_{n,srss} - \bar{x}_{r,srss}) y_{r,srss} & \text{for } i \in \bar{A} \end{cases}$$

The sequent combined estimators under the above situations are given by

$$T^c_{DP_1} = \bar{y}_{r,srss} + b_1(\bar{X} - \bar{x}_{n,srss})$$
$$T^c_{DP_2} = \bar{y}_{r,srss} + b_2(\bar{X} - \bar{x}_{r,srss})$$
$$T^c_{DP_3} = \bar{y}_{r,srss} + b_3(\bar{x}_{n,srss} - \bar{x}_{r,srss})$$

Following Sohail *et al.* ([32]), one may envisage a combined class of ratio type imputation methods under $SRSS$ for the imputation of missing values as
Situation I

$$y^c_{.iS_1} = \begin{cases} y_i & \text{for } i \in A \\ \frac{n}{n-r}\left[\bar{y}_{r,srss}\left(\frac{\bar{X}}{\bar{x}_{n,srss}}\right)^{\beta_1} - \bar{y}_{r,srss}\right] & \text{for } i \in \bar{A} \end{cases}$$

$$y^c_{.iS_4} = \begin{cases} y_i & \text{for } i \in A \\ \frac{n}{n-r}\left[\bar{y}_{r,srss}\left(\frac{\bar{X}}{\beta_4 \bar{x}_{n,srss} + (1-\beta_4)\bar{X}}\right) - \bar{y}_{r,srss}\right] & \text{for } i \in \bar{A} \end{cases}$$

Situation II

$$y^c_{.iS_2} = \begin{cases} y_i & \text{for } i \in A \\ \frac{n}{n-r}\left[\bar{y}_{r,srss}\left(\frac{\bar{X}}{\bar{x}_{r,srss}}\right)^{\beta_2} - \bar{y}_{r,srss}\right] & \text{for } i \in \bar{A} \end{cases}$$

$$y^c_{.iS_5} = \begin{cases} y_i & \text{for } i \in A \\ \frac{n}{n-r}\left[\bar{y}_{r,srss}\left(\frac{\bar{X}}{\beta_5 \bar{x}_{r,srss} + (1-\beta_5)\bar{X}}\right) - \bar{y}_{r,srss}\right] & \text{for } i \in \bar{A} \end{cases}$$

Situation III

$$y^c_{.iS_3} = \begin{cases} y_i & \text{for } i \in A \\ \frac{n}{n-r}\left[\bar{y}_{r,srss}\left(\frac{\bar{x}_{n,srss}}{\bar{x}_{r,srss}}\right)^{\beta_3} - \bar{y}_{r,srss}\right] & \text{for } i \in \bar{A} \end{cases}$$

$$y^c_{.iS_6} = \begin{cases} y_i & \text{for } i \in A \\ \frac{n}{n-r}\left[\bar{y}_{r,srss}\left(\frac{\bar{X}}{\beta_6\bar{x}_{r,srss}+(1-\beta_6)\bar{x}_{n,srss}}\right) - \bar{y}_{r,srss}\right] & \text{for } i \in \bar{A} \end{cases}$$

The sequent estimators are given by

$$T^c_{S_1} = \bar{y}_{r,srss}\left(\frac{\bar{X}}{\bar{x}_{n,srss}}\right)^{\beta_1}$$

$$T^c_{S_2} = \bar{y}_{r,srss}\left(\frac{\bar{X}}{\bar{x}_{r,srss}}\right)^{\beta_2}$$

$$T^c_{S_3} = \bar{y}_{r,srss}\left(\frac{\bar{x}_{n,srss}}{\bar{x}_{r,srss}}\right)^{\beta_3}$$

$$T^c_{S_4} = \bar{y}_{r,srss}\left(\frac{\bar{X}}{\beta_4\bar{x}_{n,srss} + (1-\beta_4)\bar{X}}\right)$$

$$T^c_{S_5} = \bar{y}_{r,srss}\left(\frac{\bar{X}}{\beta_5\bar{x}_{r,srss} + (1-\beta_5)\bar{X}}\right)$$

$$T^c_{S_6} = \bar{y}_{r,srss}\left(\frac{\bar{X}}{\beta_6\bar{x}_{r,srss} + (1-\beta_6)\bar{x}_{n,srss}}\right)$$

where $\beta_i;\ \ i = 1, 2, ..., 6$ are suitably chosen optimizing scalars.
Appendix A of supplementary file contains the minimum mean square error (MSE) of the sequent estimators consisting of different imputation methods.

3.2. Separate imputation methods

The separate mean method of imputation under $SRSS$ is given by

$$y^s_{.i_m} = \begin{cases} y_i & \text{for } i \in A_h \\ \bar{y}_{r,h[rss]} & \text{for } i \in \bar{A}_h \end{cases}$$

The sequent estimator is given by

$$T^s_m = \sum_{h=1}^{L} W_h \bar{y}_{r,h[rss]}$$

where $\bar{y}_{r,h[rss]} = \frac{1}{m_h k}\sum_{i=1}^{m_h}\sum_{j=1}^{k} y_{h[i]j}$ is the ranked set sample mean of study variable in stratum h.
The separate imputation methods are categorized into three situations under the

availability of auxiliary informations.
Situation I: When $\bar{X}$ is known and $\bar{x}_{n,h(rss)}$ is utilized.
Situation II: When $\bar{X}$ is known and $\bar{x}_{r,h(rss)}$ is utilized.
Situation III: When $\bar{X}$ is unknown and $\bar{x}_{n,h(rss)}$, $\bar{x}_{r,h(rss)}$ are utilized.
The classical separate ratio type imputation method is described under $SRSS$ as
Situation I

$$y^s_{.i_{R_1}} = \begin{cases} y_i & \text{for } i \in A_h \\ \frac{1}{n-r}\left[n\bar{y}_{r,h[rss]}\left(\frac{\bar{X}_h}{\bar{x}_{n,h(rss)}}\right) - r\bar{y}_{r,h[rss]}\right] & \text{for } i \in \bar{A}_h \end{cases}$$

Situation II

$$y^s_{.i_{R_2}} = \begin{cases} y_i & \text{for } i \in A_h \\ \frac{1}{n-r}\left[n\bar{y}_{r,h[rss]}\left(\frac{\bar{X}_h}{\bar{x}_{r,h(rss)}}\right) - r\bar{y}_{r,h[rss]}\right] & \text{for } i \in \bar{A}_h \end{cases}$$

Situation III

$$y^s_{.i_{R_3}} = \begin{cases} y_i & \text{for } i \in A_h \\ \frac{1}{n-r}\left[n\bar{y}_{r,h[rss]}\left(\frac{\bar{x}_{n,h(rss)}}{\bar{x}_{r,h(rss)}}\right) - r\bar{y}_{r,h[rss]}\right] & \text{for } i \in \bar{A}_h \end{cases}$$

The sequent estimators are given by

$$T^s_{R_1} = \sum_{h=1}^{L} W_h\left[\bar{y}_{r,h[rss]}\left(\frac{\bar{X}_h}{\bar{x}_{n,h(rss)}}\right)\right]$$

$$T^s_{R_2} = \sum_{h=1}^{L} W_h\left[\bar{y}_{r,h[rss]}\left(\frac{\bar{X}_h}{\bar{x}_{r,h(rss)}}\right)\right]$$

$$T^s_{R_3} = \sum_{h=1}^{L} W_h\left[\bar{y}_{r,h[rss]}\left(\frac{\bar{x}_{n,h(rss)}}{\bar{x}_{r,h(rss)}}\right)\right]$$

On the lines of Diana and Perri ([13]), we define a separate regression imputation method under $SRSS$ as
Situation I

$$y^s_{.i_{DP_1}} = \begin{cases} y_i & \text{for } i \in A_h \\ y_{r,h[rss]} + \frac{n}{n-r}b_1(\bar{X} - \bar{x}_{n,h(rss)}) & \text{for } i \in \bar{A}_h \end{cases}$$

Situation II

$$y^s_{.i_{DP_2}} = \begin{cases} y_i & \text{for } i \in A_h \\ y_{r,h[rss]} + \frac{n}{n-r}b_2(\bar{X} - \bar{x}_{r,h(rss)}) & \text{for } i \in \bar{A}_h \end{cases}$$

Situation III

$$y^s_{.i_{DP_3}} = \begin{cases} y_i & \text{for } i \in A_h \\ y_{r,h[rss]} + \frac{n}{n-r}b_3(\bar{x}_{n,h(rss)} - \bar{x}_{r,h(rss)}) & \text{for } i \in \bar{A}_h \end{cases}$$

The sequent separate estimators under the above situations are given by

$$T^s_{DP_1} = \sum_{h=1}^{L} W_h[\bar{y}_{r,h[rss]} + b_{1_h}(\bar{X}_h - \bar{x}_{n,h(rss)})]$$
$$T^s_{DP_2} = \sum_{h=1}^{L} W_h[\bar{y}_{r,h[rss]} + b_{2_h}(\bar{X}_h - \bar{x}_{r,h(rss)})]$$
$$T^s_{DP_3} = \sum_{h=1}^{L} W_h[\bar{y}_{r,h[rss]} + b_{3_h}(\bar{x}_{n,h(rss)} - \bar{x}_{r,h(rss)})]$$

Motivated by Sohail *et al.* ([32]), we define a separate class of ratio type imputation methods under $SRSS$ as
Situation I

$$y^s_{.is_1} = \begin{cases} y_i & \text{for } i \in A_h \\ \frac{1}{n-r}\left[n\bar{y}_{r,h[rss]}\left(\frac{\bar{X}_h}{\bar{x}_{n,h(rss)}}\right)^{\beta_{1_h}} - r\bar{y}_{r,h[rss]}\right] & \text{for } i \in \bar{A}_h \end{cases}$$
$$y^s_{.is_4} = \begin{cases} y_i & \text{for } i \in A_h \\ \frac{1}{n-r}\left[n\bar{y}_{r,h[rss]}\left(\frac{\bar{X}_h}{\beta_{4_h}\bar{x}_{n,h(rss)}+(1-\beta_{4_h})\bar{X}_h}\right) - r\bar{y}_{r,h[rss]}\right] & \text{for } i \in \bar{A}_h \end{cases}$$

Situation II

$$y^s_{.is_2} = \begin{cases} y_i & \text{for } i \in A_h \\ \frac{1}{n-r}\left[n\bar{y}_{r,h[rss]}\left(\frac{\bar{X}_h}{\bar{x}_{r,h(rss)}}\right)^{\beta_{2_h}} - r\bar{y}_{r,h[rss]}\right] & \text{for } i \in \bar{A}_h \end{cases}$$
$$y^s_{.is_5} = \begin{cases} y_i & \text{for } i \in A_h \\ \frac{1}{n-r}\left[n\bar{y}_{r,h[rss]}\left(\frac{\bar{X}_h}{\beta_{5_h}\bar{x}_{r,h(rss)}+(1-\beta_{5_h})\bar{X}_h}\right) - r\bar{y}_{r,h[rss]}\right] & \text{for } i \in \bar{A}_h \end{cases}$$

Situation III

$$y^s_{.is_3} = \begin{cases} y_i & \text{for } i \in A_h \\ \frac{1}{n-r}\left[n\bar{y}_{r,h[rss]}\left(\frac{\bar{x}_{n,h(rss)}}{\bar{x}_{r,h(rss)}}\right)^{\beta_{3_h}} - r\bar{y}_{r,h[rss]}\right] & \text{for } i \in \bar{A}_h \end{cases}$$
$$y^s_{.is_6} = \begin{cases} y_i & \text{for } i \in A_h \\ \frac{1}{n-r}\left[n\bar{y}_{r,h[rss]}\left(\frac{\bar{X}_h}{\beta_{6_h}\bar{x}_{r,h(rss)}+(1-\beta_{6_h})\bar{x}_{n,h(rss)}}\right) - r\bar{y}_{r,h[rss]}\right] & \text{for } i \in \bar{A}_h \end{cases}$$

The sequent estimators are given by

$$T^s_{S_1} = \sum_{h=1}^{L} W_h\bar{y}_{r,h[rss]}\left(\frac{\bar{X}_h}{\bar{x}_{n,h(rss)}}\right)^{\beta_{1_h}}$$

$$T^s_{S_2} = \sum_{h=1}^{L} W_h \bar{y}_{r,h[rss]} \left(\frac{\bar{X}_h}{\bar{x}_{r,h(rss)}}\right)^{\beta_{2_h}}$$

$$T^s_{S_3} = \sum_{h=1}^{L} W_h \bar{y}_{r,h[rss]} \left(\frac{\bar{x}_{n,h(rss)}}{\bar{x}_{r,h(rss)}}\right)^{\beta_{3_h}}$$

$$T^s_{S_4} = \sum_{h=1}^{L} W_h \bar{y}_{r,h[rss]} \left(\frac{\bar{X}_h}{\beta_{4_h}\bar{x}_{n,h(rss)} + (1-\beta_{4_h})\bar{X}_h}\right)$$

$$T^s_{S_5} = \sum_{h=1}^{L} W_h \bar{y}_{r,h[rss]} \left(\frac{\bar{X}_h}{\beta_{5_h}\bar{x}_{r,h(rss)} + (1-\beta_{5_h})\bar{X}_h}\right)$$

$$T^s_{S_6} = \sum_{h=1}^{L} W_h \bar{y}_{r,h[rss]} \left(\frac{\bar{X}_h}{\beta_{6_h}\bar{x}_{r,h(rss)} + (1-\beta_{6_h})\bar{x}_{n,h(rss)}}\right)$$

where $\beta_{i_h};\ \ i = 1, 2, ..., 6$ are suitably opted scalars.
Appendix B of supplementary file contains the minimum mean square error (MSE) of the sequent estimators consisting of different imputation methods.

4. Proposed imputation methods

The crux of this paper is binary:

1. To propose some efficient combined and separate imputation methods for the estimation of population mean $\bar{Y}$.
2. To determine the effect of the correlation coefficient, coefficient of skewness, and coefficient of kurtosis over the efficiency of the imputation procedures.

4.1. Combined imputation methods

Following Bhushan and Pandey ([4, 5]), we envisage nine new imputation methods under the three situations specified in the former section as
Situation I

$$y^c_{.i_{SA_1}} = \begin{cases} \alpha_1 y_i & \text{for } i \in A \\ \alpha_1 \bar{y}_{r,srss} + \frac{n\theta_1}{n-r}(\bar{x}_{n,srss} - \bar{X}) & \text{for } i \in \bar{A} \end{cases}$$

$$y^c_{.i_{SA_4}} = \begin{cases} y_i & \text{for } i \in A \\ \frac{1}{n-r}\left[n\alpha_4 \bar{y}_{r,srss}\left(\frac{\bar{X}}{\bar{x}_{n,srss}}\right)^{\theta_4} - r\bar{y}_{r,srss}\right] & \text{for } i \in \bar{A} \end{cases}$$

$$y^c_{.i_{SA_7}} = \begin{cases} y_i & \text{for } i \in A \\ \frac{1}{n-r}\left[n\alpha_7 \bar{y}_{r,srss}\left(\frac{\bar{X}}{\bar{X}+\theta_7(\bar{x}_{n,srss}-\bar{X})}\right) - r\bar{y}_{r,srss}\right] & \text{for } i \in \bar{A} \end{cases}$$

Situation II

$$y^c_{.i_{SA_2}} = \begin{cases} \alpha_2 y_i & \text{for } i \in A \\ \alpha_2 \bar{y}_{r,srss} + \frac{n\theta_2}{n-r}(\bar{x}_{r,srss} - \bar{X}) & \text{for } i \in \bar{A} \end{cases}$$

$$y^c_{.i_{SA_5}} = \begin{cases} y_i & \text{for } i \in A \\ \frac{1}{n-r}\left[n\alpha_5 \bar{y}_{r,srss}\left(\frac{\bar{X}}{\bar{x}_{r,srss}}\right)^{\theta_5} - r\bar{y}_{r,srss}\right] & \text{for } i \in \bar{A} \end{cases}$$

$$y^c_{.i_{SA_8}} = \begin{cases} y_i & \text{for } i \in A \\ \frac{1}{n-r}\left[n\alpha_8 \bar{y}_{r,srss}\left(\frac{\bar{X}}{\bar{X}+\theta_8(\bar{x}_{r,srss}-\bar{X})}\right) - r\bar{y}_{r,srss}\right] & \text{for } i \in \bar{A} \end{cases}$$

Situation III

$$y^c_{.i_{SA_3}} = \begin{cases} \alpha_3 y_i & \text{for } i \in A \\ \alpha_3 \bar{y}_{r,srss} + \frac{n\theta_3}{n-r}(\bar{x}_{r,srss} - \bar{x}_{n,srss}) & \text{for } i \in \bar{A} \end{cases}$$

$$y^c_{.i_{SA_6}} = \begin{cases} y_i & \text{for } i \in A \\ \frac{1}{n-r}\left[n\alpha_6 \bar{y}_{r,srss}\left(\frac{\bar{x}_{n,srss}}{\bar{x}_{r,srss}}\right)^{\theta_6} - r\bar{y}_{r,srss}\right] & \text{for } i \in \bar{A} \end{cases}$$

$$y^c_{.i_{SA_9}} = \begin{cases} y_i & \text{for } i \in A \\ \frac{1}{n-r}\left[n\alpha_9 \bar{y}_{r,srss}\left(\frac{\bar{x}_{n,srss}}{\bar{x}_{n,srss}+\theta_9(\bar{x}_{n,srss}-\bar{x}_{r,srss})}\right) - r\bar{y}_{r,srss}\right] & \text{for } i \in \bar{A} \end{cases}$$

Under the above situations, the sequent estimators are given by

$$\begin{aligned}
T^c_{SA_1} &= \alpha_1 \bar{y}_{r,srss} + \theta_1(\bar{x}_{n,srss} - \bar{X}) \\
T^c_{SA_2} &= \alpha_2 \bar{y}_{r,srss} + \theta_2(\bar{x}_{r,srss} - \bar{X}) \\
T^c_{SA_3} &= \alpha_3 \bar{y}_{r,srss} + \theta_3(\bar{x}_{r,srss} - \bar{x}_{n,srss}) \\
T^c_{SA_4} &= \alpha_4 \bar{y}_{r,srss}\left(\frac{\bar{X}}{\bar{x}_{n,srss}}\right)^{\theta_4} \\
T^c_{SA_5} &= \alpha_5 \bar{y}_{r,srss}\left(\frac{\bar{X}}{\bar{x}_{r,srss}}\right)^{\theta_5} \\
T^c_{SA_6} &= \alpha_6 \bar{y}_{r,srss}\left(\frac{\bar{x}_{n,srss}}{\bar{x}_{r,srss}}\right)^{\theta_6} \\
T^c_{SA_7} &= \alpha_7 \bar{y}_{r,srss}\left[\frac{\bar{X}}{\bar{X}+\theta_7(\bar{x}_{n,srss}-\bar{X})}\right] \\
T^c_{SA_8} &= \alpha_8 \bar{y}_{r,srss}\left[\frac{\bar{X}}{\bar{X}+\theta_8(\bar{x}_{r,srss}-\bar{X})}\right] \\
T^c_{SA_9} &= \alpha_9 \bar{y}_{r,srss}\left[\frac{\bar{x}_{n,srss}}{\bar{x}_{n,srss}+\theta_9(\bar{x}_{r,srss}-\bar{x}_{n,srss})}\right]
\end{aligned}$$

where $\alpha_1, \alpha_2, ..., \alpha_9$ and $\theta_1, \theta_2, ..., \theta_9$ are the suitably chosen scalars.

Theorem 4.1. *The MSE of the sequent estimators consisting of the proposed imputation methods is given by*

$$MSE(T^c_{SA_1}) = (\alpha_1 - 1)^2\bar{Y}^2 + \alpha_1^2\bar{Y}^2 I_0^* + \theta_1^2\bar{X}^2 I_1 + 2\alpha_1\theta_1\bar{X}\bar{Y}I_{01}$$

$$MSE(T^c_{SA_2}) = (\alpha_2-1)^2\bar{Y}^2 + \alpha_2^2\bar{Y}^2 I_0^* + \theta_2^2\bar{X}^2 I_1^* + 2\alpha_2\theta_2\bar{X}\bar{Y}I_{01}^*$$

$$MSE(T^c_{SA_3}) = \begin{bmatrix} (\alpha_3-1)^2\bar{Y}^2 + \alpha_3^2\bar{Y}^2 I_0^* + \theta_3^2\bar{X}^2\left\{I_1^*-I_1\right\} \\ +2\alpha_3\theta_3\bar{X}\bar{Y}\left\{I_{01}^*-I_{01}\right\} \end{bmatrix}$$

$$MSE(T^c_{SA_4}) = \bar{Y}^2 \begin{bmatrix} 1+\alpha_4^2\left\{1+I_0^*+\theta_4(2\theta_4+1)I_1-4\theta_4 I_{01}\right\} \\ -2\alpha_4\left\{1-\theta_4 I_{01}+\frac{\theta_4(\theta_4+1)}{2}I_1\right\} \end{bmatrix}$$

$$MSE(T^c_{SA_5}) = \bar{Y}^2 \begin{bmatrix} 1+\alpha_5^2\left\{1+I_0^*+\theta_5(2\theta_5+1)I_1^*-4\theta_5 I_{01}^*\right\} \\ -2\alpha_5\left\{1-\theta_5 I_{01}^*+\frac{\theta_5(\theta_5+1)}{2}I_1^*\right\} \end{bmatrix}$$

$$MSE(T^c_{SA_6}) = \bar{Y}^2 \begin{bmatrix} 1+\alpha_6^2\left\{1+I_0^*+\theta_6(2\theta_6+1)(I_1^*-I_1)-4\theta_6(I_{01}^*-I_{01})\right\} \\ -2\alpha_6\left\{1-\theta_6(I_{01}^*-I_{01})+\frac{\theta_6(\theta_6+1)}{2}(I_1^*-I_1)\right\} \end{bmatrix}$$

$$MSE(T^c_{SA_7}) = \bar{Y}^2 \begin{bmatrix} 1+\alpha_7^2\left\{1+I_0^*+3\theta_7^2 I_1-4\theta_7 I_{01}\right\} \\ -2\alpha_7\left\{1+\theta_7^2 I_1-\theta_7 I_{01}\right\} \end{bmatrix}$$

$$MSE(T^c_{SA_8}) = \bar{Y}^2 \begin{bmatrix} 1+\alpha_8^2\left\{1+I_0^*+3\theta_8^2 I_1^*-4\theta_8 I_{01}^*\right\} \\ -2\alpha_8\left\{1+\theta_8^2 I_1^*-\theta_8 I_{01}^*\right\} \end{bmatrix}$$

$$MSE(T^c_{SA_9}) = \bar{Y}^2 \begin{bmatrix} 1+\alpha_9^2\left\{1+I_0^*+3\theta_9^2(I_1^*-I_1)-4\theta_9(I_{01}^*-I_{01})\right\} \\ -2\alpha_9\left\{1+\theta_9^2(I_1^*-I_1)-\theta_9(I_{01}^*-I_{01})\right\} \end{bmatrix}$$

Proof: Appendix C of supplementary file contains a summary of the derivations. The derivations can easily be done using Taylor series expansion. □

Theorem 4.2. *The minimum MSE of the sequent estimators consisting of the proposed imputation methods are*

$$minMSE(T^c_{SA_i}) = \bar{Y}^2(1-\alpha_{i(opt)}) = \bar{Y}^2\left(1-\frac{A_i^2}{B_i}\right);\ i=1,2,3 \tag{4.1}$$

$$minMSE(T^c_{SA_j}) = \bar{Y}^2\left(1-\frac{A_j^2}{B_j}\right);\ j=4,5,6 \tag{4.2}$$

$$minMSE(T^c_{SA_k}) = \bar{Y}^2\left(1-\frac{A_k^2}{B_k}\right);\ k=7,8,9 \tag{4.3}$$

Proof: Appendix C of supplementary file contains a summary of the derivations. □

Corollary 4.1. *The proposed sequent estimators $T^c_{SA_i},\ i=1,2,3$ dominate the proposed sequent estimators $T^c_{SA_j},\ j=4,5,6,$ iff*

$$\alpha_{i(opt)} > \frac{A_j^2}{B_j} \tag{4.4}$$

and contrariwise. Otherwise, both are equally efficient when the equality holds in (4.4).

Proof: By comparing the minimum MSEs of the proposed estimators from (4.1) and (4.2), we get (4.4). □

Corollary 4.2. *The proposed sequent estimators $T^c_{SA_i}$, $i = 1, 2, 3$ dominate the proposed sequent estimators $T^c_{SA_k}$, $k = 7, 8, 9$, iff*

$$\alpha_{i(opt)} > \frac{A_k^2}{B_k} \tag{4.5}$$

and contrariwise. Otherwise, both are equally efficient when the equality holds in (4.5).

Proof: On comparing the minimum MSEs of the proposed estimators from (4.1) and (4.3), we get (4.5). □

Corollary 4.3. *The proposed sequent estimators $T^c_{SA_j}$, $i = 4, 5, 6$ dominate the proposed sequent estimators $T^c_{SA_k}$, $k = 7, 8, 9$, iff*

$$\frac{A_j^2}{B_j} > \frac{A_k^2}{B_k} \tag{4.6}$$

and contrariwise. Otherwise, both are equally efficient when the equality holds in (4.6).

Proof: On comparing the minimum MSEs of the proposed estimators from (4.2) and (4.3), we get (4.6). □

The only way to determine if (4.4), (4.5), and (4.6) are true in practise is through the computational analysis done in Section 6.

4.2. Separate imputation methods

On the lines of Bhushan and Pandey ([4, 5]), we suggest nine new separate imputation methods under the three situations discussed in the earlier section as *Situation I*

$$y^s_{.i_{SA_1}} = \begin{cases} \alpha_{1_h} y_i & \text{for } i \in A_h \\ \alpha_{1_h} \bar{y}_{r,h[rss]} + \frac{n\theta_{1_h}}{n-r}(\bar{x}_{n,h(rss)} - \bar{X}_h) & \text{for } i \in \bar{A}_h \end{cases}$$

$$y^s_{.i_{SA_4}} = \begin{cases} y_i & \text{for } i \in A_h \\ \frac{1}{n-r}\left[n\alpha_{4_h}\bar{y}_{r,h[rss]}\left(\frac{\bar{X}_h}{\bar{x}_{n,h(rss)}}\right)^{\theta_{4_h}} - r\bar{y}_{r,h[rss]}\right] & \text{for } i \in \bar{A}_h \end{cases}$$

$$y^s_{.i_{SA_7}} = \begin{cases} y_i & \text{for } i \in A_h \\ \frac{1}{n-r}\left[n\alpha_{7_h}\bar{y}_{r,h[rss]}\left(\frac{\bar{X}_h}{\bar{X}_h + \theta_{7_h}(\bar{x}_{n,h(rss)} - \bar{X}_h)}\right) - r\bar{y}_{r,h[rss]}\right] & \text{for } i \in \bar{A}_h \end{cases}$$

Situation II

$$y^s_{.i_{SA_2}} = \begin{cases} \alpha_{2_h} y_i & \text{for } i \in A_h \\ \alpha_{2_h} \bar{y}_{r,h[rss]} + \frac{n\theta_{2_h}}{n-r}(\bar{x}_{r,h(rss)} - \bar{X}_h) & \text{for } i \in \bar{A}_h \end{cases}$$

$$y^s_{.i_{SA_5}} = \begin{cases} y_i & \text{for } i \in A_h \\ \frac{1}{n-r}\left[n\alpha_{5_h} \bar{y}_{r,h[rss]} \left(\frac{\bar{X}_h}{\bar{x}_{r,h(rss)}}\right)^{\theta_{5_h}} - r\bar{y}_{r,h[rss]}\right] & \text{for } i \in \bar{A}_h \end{cases}$$

$$y^s_{.i_{SA_8}} = \begin{cases} y_i & \text{for } i \in A_h \\ \frac{1}{n-r}\left[n\alpha_{8_h} \bar{y}_{r,h[rss]} \left(\frac{\bar{X}_h}{\bar{X}+\theta_{8_h}(\bar{x}_{r,h(rss)}-\bar{X}_h)}\right) - r\bar{y}_{r,h[rss]}\right] & \text{for } i \in \bar{A}_h \end{cases}$$

Situation III

$$y^s_{.i_{SA_3}} = \begin{cases} \alpha_{3_h} y_i & \text{for } i \in A_h \\ \alpha_{3_h} \bar{y}_{r,h[rss]} + \frac{n\theta_{3_h}}{n-r}(\bar{x}_{r,h(rss)} - \bar{x}_{n,h(rss)}) & \text{for } i \in \bar{A}_h \end{cases}$$

$$y^s_{.i_{SA_6}} = \begin{cases} y_i & \text{for } i \in A_h \\ \frac{1}{n-r}\left[n\alpha_{6_h} \bar{y}_{r,h[rss]} \left(\frac{\bar{x}_{n,h(rss)}}{\bar{x}_{r,h(rss)}}\right)^{\theta_{6_h}} - r\bar{y}_{r,h[rss]}\right] & \text{for } i \in \bar{A}_h \end{cases}$$

$$y^s_{.i_{SA_9}} = \begin{cases} y_i & \text{for } i \in A_h \\ \frac{1}{n-r}\left[n\alpha_{9_h} \bar{y}_{r,h[rss]} \left(\frac{\bar{x}_{n,h(rss)}}{\bar{x}_{n,h(rss)}+\theta_{9_h}(\bar{x}_{n,h(rss)}-\bar{x}_{r,h(rss)})}\right) - r\bar{y}_{r,h[rss]}\right] & \text{for } i \in \bar{A}_h \end{cases}$$

The sequent estimators consisting of the above imputation methods are

$$T^s_{SA_1} = \sum_{h=1}^{L} W_h[\alpha_{1_h} \bar{y}_{r,h[rss]} + \theta_{1_h}(\bar{x}_{n,h(rss)} - \bar{X}_h)]$$

$$T^s_{SA_2} = \sum_{h=1}^{L} W_h[\alpha_{2_h} \bar{y}_{r,h[rss]} + \theta_{2_h}(\bar{x}_{r,h(rss)} - \bar{X}_h)]$$

$$T^s_{SA_3} = \sum_{h=1}^{L} W_h[\alpha_{3_h} \bar{y}_{r,h[rss]} + \theta_{3_h}(\bar{x}_{r,h(rss)} - \bar{x}_{n,h(rss)})]$$

$$T^s_{SA_4} = \sum_{h=1}^{L} W_h \alpha_{4_h} \bar{y}_{r,h[rss]} \left(\frac{\bar{X}_h}{\bar{x}_{n,h(rss)}}\right)^{\theta_{4_h}}$$

$$T^s_{SA_5} = \sum_{h=1}^{L} W_h \alpha_{5_h} \bar{y}_{r,h[rss]} \left(\frac{\bar{X}_h}{\bar{x}_{r,h(rss)}}\right)^{\theta_{5_h}}$$

$$T^s_{SA_6} = \sum_{h=1}^{L} W_h \alpha_{6_h} \bar{y}_{r,h[rss]} \left(\frac{\bar{x}_{n,h(rss)}}{\bar{x}_{r,h(rss)}}\right)^{\theta_{6_h}}$$

$$T^s_{SA_7} = \sum_{h=1}^{L} W_h \alpha_{7_h} \bar{y}_{r,h[rss]} \left[\frac{\bar{X}_h}{\bar{X}_h + \theta_{7_h}(\bar{x}_{n,h(rss)} - \bar{X}_h)}\right]$$

$$T^s_{SA_8} = \sum_{h=1}^{L} W_h \alpha_{8_h} \bar{y}_{r,h[rss]} \left[\frac{\bar{X}_h}{\bar{X}_h + \theta_{8_h}(\bar{x}_{r,h(rss)} - \bar{X}_h)} \right]$$

$$T^s_{SA_9} = \sum_{h=1}^{L} W_h \alpha_{9_h} \bar{y}_{r,h[rss]} \left[\frac{\bar{x}_{n,h(rss)}}{\bar{x}_{n,h(rss)} + \theta_{9_h}(\bar{x}_{r,h(rss)} - \bar{x}_{n,h(rss)})} \right]$$

where $\alpha_{1_h}, \alpha_{2_h}, ..., \alpha_{9_h}$ and $\theta_{1_h}, \theta_{2_h}, ..., \theta_{9_h}$ are suitably chosen scalars.

Theorem 4.3. *The MSE of the sequent estimators consisting of the proposed imputation methods is given by*

$$MSE(T^s_{SA_1}) = \sum_{h=1}^{L} W_h^2 \left[(\alpha_{1_h} - 1)^2 \bar{Y}_h^2 + \alpha_{1_h}^2 \bar{Y}_h^2 J_0^* + \theta_{1_h}^2 \bar{X}_h^2 J_1 + 2\alpha_{1_h}\theta_{1_h}\bar{X}_h\bar{Y}_h J_{01} \right]$$

$$MSE(T^s_{SA_2}) = \sum_{h=1}^{L} W_h^2 \left[(\alpha_{2_h} - 1)^2 \bar{Y}_h^2 + \alpha_{2_h}^2 \bar{Y}_h^2 J_0^* + \theta_{2_h}^2 \bar{X}_h^2 J_1^* + 2\alpha_{2_h}\theta_{2_h}\bar{X}_h\bar{Y}_h J_{01}^* \right]$$

$$MSE(T^s_{SA_3}) = \sum_{h=1}^{L} W_h^2 \left[\begin{array}{l} (\alpha_{3_h} - 1)^2 \bar{Y}_h^2 + \alpha_{3_h}^2 \bar{Y}_h^2 J_0^* + \theta_{3_h}^2 \bar{X}_h^2 \{J_1^* - J_1\} \\ +2\alpha_{3_h}\theta_{3_h}\bar{X}_h\bar{Y}_h \{J_{01}^* - J_{01}\} \end{array} \right]$$

$$MSE(T^s_{SA_4}) = \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left[\begin{array}{l} 1 + \alpha_{4_h}^2 \{ 1 + J_0^* + \theta_{4_h}(2\theta_{4_h} + 1)J_1 - 4\theta_{4_h} J_{01} \} \\ -2\alpha_{4_h} \left\{ 1 - \theta_{4_h} J_{01} + \frac{\theta_{4_h}(\theta_{4_h}+1)}{2} J_1 \right\} \end{array} \right]$$

$$MSE(T^s_{SA_5}) = \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left[\begin{array}{l} 1 + \alpha_{5_h}^2 \{ 1 + J_0^* + \theta_{5_h}(2\theta_{5_h} + 1)J_1^* - 4\theta_{5_h} J_{01}^* \} \\ -2\alpha_{5_h} \left\{ 1 - \theta_{5_h} J_{01}^* + \frac{\theta_{5_h}(\theta_{5_h}+1)}{2} J_1^* \right\} \end{array} \right]$$

$$MSE(T^s_{SA_6}) = \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left[\begin{array}{l} 1 + \alpha_{6_h}^2 \{ 1 + J_0^* + \theta_{6_h}(2\theta_{6_h} + 1)(J_1^* - J_1) - 4\theta_{6_h}(J_{01}^* - J_{01}) \} \\ -2\alpha_{6_h} \left\{ 1 - \theta_{6_h}(J_{01}^* - J_{01}) + \frac{\theta_{6_h}(\theta_{6_h}+1)}{2}(J_1^* - J_1) \right\} \end{array} \right]$$

$$MSE(T^s_{SA_7}) = \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left[\begin{array}{l} 1 + \alpha_{7_h}^2 \{ 1 + J_0^* + 3\theta_{7_h}^2 J_1 - 4\theta_{7_h} J_{01} \} \\ -2\alpha_{7_h} \{1 + \theta_{7_h}^2 J_1 - \theta_{7_h} J_{01}\} \end{array} \right]$$

$$MSE(T^s_{SA_8}) = \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left[\begin{array}{l} 1 + \alpha_{8_h}^2 \{ 1 + J_0^* + 3\theta_{8_h}^2 J_1^* - 4\theta_{8_h} J_{01}^* \} \\ -2\alpha_{8_h} \{1 + \theta_{8_h}^2 J_1^* - \theta_{8_h} J_{01}^*\} \end{array} \right]$$

$$MSE(T^s_{SA_9}) = \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left[\begin{array}{l} 1 + \alpha_{9_h}^2 \{ 1 + J_0^* + 3\theta_{9_h}^2(J_1^* - J_1) - 4\theta_{9_h}(J_{01}^* - J_{01}) \} \\ -2\alpha_{9_h} \{ 1 + \theta_{9_h}^2(J_1^* - J_1) - \theta_{9_h}(J_{01}^* - J_{01}) \} \end{array} \right]$$

Proof: Appendix C of supplementary file contains a summary of the derivations. The derivations can easily be done using Taylor series expansion. □

Theorem 4.4. *The minimum MSE of the sequent estimators consisting* *of the proposed imputation methods is given by*

(4.7)
$$minMSE(T^s_{SA_i}) = \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 (1-\alpha_{i_h(opt)}) = \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left(1-\frac{A_{i_h}^2}{B_{i_h}}\right);\ i=1,2,3$$

(4.8)
$$minMSE(T^s_{SA_j}) = \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left(1-\frac{A_{j_h}^2}{B_{j_h}}\right);\ j=4,5,6$$

(4.9)
$$minMSE(T^s_{SA_k}) = \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left(1-\frac{A_{k_h}^2}{B_{k_h}}\right);\ k=7,8,9$$

Proof: Appendix C of supplementary file contains a summary of the derivations. □

Corollary 4.4. *The proposed sequent estimators $T^s_{SA_i},\ i=1,2,3$ dominate the proposed sequent estimators $T^s_{SA_j},\ j=4,5,6$, iff*

$$\sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \alpha_{i_h(opt)} > \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left(\frac{A_{i_h}^2}{B_{i_h}}\right) \tag{4.10}$$

and contrariwise. Otherwise, both are equally efficient when the equality holds in (4.10).

Proof: On comparing the minimum MSEs of the proposed estimators from (4.7) and (4.8), we get (4.10). □

Corollary 4.5. *The proposed sequent estimators $T^s_{SA_i},\ i=1,2,3$ dominate the proposed sequent estimators $T^s_{SA_k},\ k=4,5,6$, iff*

$$\sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \alpha_{i_h(opt)} > \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left(\frac{A_{k_h}^2}{B_{k_h}}\right) \tag{4.11}$$

and contrariwise. Otherwise, both are equally efficient when the equality holds in (4.11).

Proof: By comparing the minimum MSEs of the proposed estimators from (4.7) and (4.9), we get (4.11). □

Corollary 4.6. *The proposed sequent estimators $T^s_{SA_j},\ j = 4, 5, 6$ dominate the proposed sequent estimators $T^s_{SA_k},\ k = 7, 8, 9$, iff*

$$\sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left(\frac{A_{j_h}^2}{B_{j_h}}\right) > \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left(\frac{A_{k_h}^2}{B_{k_h}}\right) \tag{4.12}$$

and contrariwise. Otherwise, both are equally efficient when the equality holds in (4.12).

Proof: By comparing the minimum MSEs of the proposed estimators from (4.8) and (4.9), we get (4.12). □

The only way to determine if (4.10), (4.11), and (4.12) are true in practise is through the computational analysis done in Section 6.

5. Optimality conditions

In this section, we obtain the optimality conditions under two heads, namely, optimality conditions for combined imputation methods and the optimality conditions for separate imputation methods.

5.1. Optimality conditions for the combined imputation methods

By comparing the minimum MSE of the suggested combined imputation methods $y^c_{.i_{SA_i}},\ i = 1, 2, ..., 9$ from (4.1) and (4.2) with the minimum MSE of the other existing combined imputation methods from $(A.1)$, $(A.2)$, $(A.3)$, $(A.4)$, $(A.8)$, $(A.9)$, $(A.10)$, $(A.14)$, $(A.15)$, and $(A.16)$, respectively, given in Appendix A of supplementary file, we get the following optimality conditions.

$$MSE(T^c_m) > MSE(T^c_{SA_i}) \implies \frac{A_i^2}{B_i} > 1 - I_0^*$$

$$MSE(T^c_{R_1}) > MSE(T^c_{SA_i}) \implies \frac{A_i^2}{B_i} > 1 - I_0^* - I_1 + 2I_{01}$$

$$MSE(T^c_{R_2}) > MSE(T^c_{SA_i}) \implies \frac{A_i^2}{B_i} > 1 - I_0^* - I_1^* + 2I_{01}^*$$

$$MSE(T^c_{R_3}) > MSE(T^c_{SA_i}) \implies \frac{A_i^2}{B_i} > 1 - I_0^* - I_1 - I_1^* + 2I_{01}^*$$

$$MSE(T^c_{DP_1}) > MSE(T^c_{SA_i}) \implies \frac{A_i^2}{B_i} > 1 - I_0^* + \frac{I_{01}^2}{I_1}$$

$$MSE(T^c_{DP_2}) > MSE(T^c_{SA_i}) \implies \frac{A_i^2}{B_i} > 1 - I_0^* + \frac{{I_{01}^*}^2}{I_1^*}$$
$$MSE(T^c_{DP_3}) > MSE(T^c_{SA_i}) \implies \frac{A_i^2}{B_i} > 1 - I_0^* + \frac{(I_{01}^* - I_{01})^2}{(I_1^* - I_1)}$$
$$MSE(T^c_{S_1}) > MSE(T^c_{SA_i}) \implies \frac{A_i^2}{B_i} > 1 - I_0^* + \frac{I_{01}^2}{I_1}$$
$$MSE(T^c_{S_2}) > MSE(T^c_{SA_i}) \implies \frac{A_i^2}{B_i} > 1 - I_0^* + \frac{{I_{01}^*}^2}{I_1^*}$$
$$MSE(T^c_{S_3}) > MSE(T^c_{SA_i}) \implies \frac{A_i^2}{B_i} > 1 - I_0^* + \frac{(I_{01}^* - I_{01})^2}{(I_1^* - I_1)}$$

The optimality of the suggested combined imputation methods can be justified under the above conditions.

5.2. Optimality conditions for the separate imputation methods

By comparing the minimum MSE of the proposed imputation methods $y^s_{.i_{SA_i}}$, $i = 1, 2, ..., 9$ given in (4.7) and (4.8) with the minimum MSE of the other existing imputation methods given in $(B.17)$, $(B.18)$, $(B.19)$, $(B.20)$, $(B.24)$, $(B.25)$, $(B.26)$, $(B.30)$, $(B.31)$, and $(B.32)$, respectively, given in Appendix B of supplementary file, we get the following optimality conditions.

$$MSE(T^s_m) > MSE(T^s_{SA_i}) \implies \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left(1 - \frac{A_{i_h}^2}{B_{i_h}}\right) < \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 J_1^*$$
$$MSE(T^s_{R_1}) > MSE(T^s_{SA_i}) \implies \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left(1 - \frac{A_{i_h}^2}{B_{i_h}}\right) < \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left[J_0^* + J_1 - 2J_{01}\right]$$
$$MSE(T^s_{R_2}) > MSE(T^s_{SA_i}) \implies \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left(1 - \frac{A_{i_h}^2}{B_{i_h}}\right) < \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left[J_0^* + J_1^* - 2J_{01}^*\right]$$
$$MSE(T^s_{R_3}) > MSE(T^s_{SA_i}) \implies \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left(1 - \frac{A_{i_h}^2}{B_{i_h}}\right) < \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left[J_0^* + J_1^* + J_1 - 2J_{01}^*\right]$$
$$MSE(T^s_{DP_1}) > MSE(T^s_{SA_i}) \implies \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left(1 - \frac{A_{i_h}^2}{B_{i_h}}\right) < \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left[J_0^* - \frac{J_{01}^2}{J_1}\right]$$
$$MSE(T^s_{DP_2}) > MSE(T^s_{SA_i}) \implies \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left(1 - \frac{A_{i_h}^2}{B_{i_h}}\right) < \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left[J_0^* - \frac{{J_{01}^*}^2}{J_1^*}\right]$$
$$MSE(T^s_{DP_3}) > MSE(T^s_{SA_i}) \implies \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left(1 - \frac{A_{i_h}^2}{B_{i_h}}\right) < \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left[J_0^* - \frac{(J_{01}^* - J_{01})^2}{\left(J_1^* - J_1\right)}\right]$$

$$MSE(T^s_{S_1}) > MSE(T^s_{SA_i}) \implies \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left(1 - \frac{A^2_{i_h}}{B_{i_h}}\right) < \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left[J_0^* - \frac{J_{01}^2}{J_1}\right]$$

$$MSE(T^s_{S_2}) > MSE(T^s_{SA_i}) \implies \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left(1 - \frac{A^2_{i_h}}{B_{i_h}}\right) < \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left[J_0^* - \frac{{J_{01}^*}^2}{J_1^*}\right]$$

$$MSE(T^s_{S_3}) > MSE(T^s_{SA_i}) \implies \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left(1 - \frac{A^2_{i_h}}{B_{i_h}}\right) < \sum_{h=1}^{L} W_h^2 \bar{Y}_h^2 \left[J_0^* - \frac{\left(J_{01}^* - J_{01}\right)^2}{\left(J_1^* - J_1\right)}\right]$$

Under the above conditions, the optimality of the proposed separate imputation methods can be ascertained.

5.3. Comparison of proposed combined and separate imputation methods

By comparing the minimum MSE of the proposed combined and separate classes of imputation methods from (4.1), (4.2) and (4.7), (4.8), we get

(5.1)

$$minMSE(T^c_{SA_i}) - minMSE(T^s_{SA_i}) = \sum_{h=1}^{L}\left[(\bar{Y}^2 - W_h^2\bar{Y}_h^2) - \left(\bar{Y}^2\frac{A_i^2}{B_i} - W_h^2\bar{Y}_h^2\frac{A^2_{i_h}}{B_{i_h}}\right)\right]$$

If the sequent estimators are conclusive and the relationship between auxiliary and study variables within each stratum is a straight line passing through the origine, then the last term of (5.1) is miniscule and it vanished.
In addition, except R_h becomes invariant from stratum to stratum, the separate estimators perform better in each stratum provided the sample in each stratum is to be sufficiently large so that the approximate formula for $MSE(T^s_{SA_i}),\ i = 1, 2, ..., 9$ is valid and the cumulative bias that can affect the proposed estimators is negligible, whereas the proposed combined estimators are to be highly advocated with only a small sample in each stratum (see, Cochran ([11])).

6. Simulation Study

To highlight the properties and to access the performance of the proposed imputation methods, motivated by Singh and Horn ([29]), simulations were carried out over two artificially generated asymmetric populations such as gamma and exponential of size $N = 2100$ units each with variables X and Y whose values are given by

$$y_i = 8.2 + \sqrt{(1-\rho_{xy}^2)}\ y_i^* + \rho_{xy}\left(\frac{S_y}{S_x}\right)x_i^*$$

$$x_i = 4.2 + x_i^*$$

where x_i^* and y_i^* are independent variates of proportional distribution. Each population is divided into three equal strata and a stratified ranked set sample of size 9 with set size 3 and number of cycles 3 is drawn from each stratum with the help of the methodology described in Section 2. With 10000 iterations, the percent relative efficiency (PRE) of the sequent estimators with respect to the conventional mean estimator was obtained as

$$PRE = \frac{\frac{1}{10000}\sum_{i=1}^{10000}(T_m - \bar{Y})^2}{\frac{1}{10000}\sum_{i=1}^{10000}(T^* - \bar{Y})^2} \times 100$$

where T^* is the existing and proposed combined and separate class of estimators. The findings of the simulation are disclosed from Table 1 to Table 4 through PRE for reasonably chosen values of correlation coefficient $\rho_{xy} = 0.6, 0.7, 0.8, 0.9$ and fair choice of response probability $P = 0.4, 0.6$.
From Table 1 to Table 4, consisting of the simulation results of two asymmetric populations, namely, gamma and exponential, we have seen that the proposed combined and separate imputation methods $y^c_{.i_{SA_j}}$ and $y^s_{.i_{SA_j}}$, $j = 1, 2, ..., 9$ dominate the other existing imputation methods for reasonably chosen values of the correlation coefficient. We have also seen that the proposed combined and separate imputation methods $y^c_{.i_{SA_j}}$ and $y^s_{.i_{SA_j}}$, $j = 4, 5, 6$ perform better among the proposed class of imputation methods under situations I, II and III. Moreover, it is also seen that the PRE of the proposed imputation methods under situations I, II and III in both populations decreases with the increase in asymmetry.

7. Conclusion

This paper is the outset to suggest some combined and separate classes of imputation methods along with their properties for the estimation of population mean in the presence of missing data using $SRSS$. The theoretical conditions are derived under which the proposed combined and separate classes of imputation methods are justified. In order to enhance the theoretical findings and to determine the effect of skewness and kurtosis over PRE, a simulation study is accomplished on two asymmetric populations viz. gamma and exponential with reasonable choice of correlation coefficient ρ_{xy} and probability of non responding units P. It is noticed from the perusal of theoretical and simulation results that:

1. The proposed combined and separate of imputation methods $y^c_{.i_{SA_j}}$ and $y^s_{.i_{SA_j}}$, $j = 1, 2, ..., 9$ always perform better than the combined and separate mean imputation method $y^c_{.i_m}$ and $y^s_{.i_m}$, ratio imputation methods $y^c_{.i_{R_j}}$ and $y^s_{.i_{R_j}}$, $j = 1, 2, 3$ and their own conventional counterparts for different values of correlation coefficient ρ_{xy}, coefficient of skewness β_1 and coefficient of kurtosis β_2.

2. The proposed combined and separate imputation methods $y^c_{.i_{SA_j}}$ and $y^s_{.i_{SA_j}}$,

Table 1: *PRE* of proposed combined estimators at P=0.4

ρ_{xy}	0.6	0.7	0.8	0.9
T_m^c	100	100	100	100
$x^* \sim \Gamma(0.5, 1.5)$				
$y^* \sim \Gamma(1.5, 2)$				
Skewness of y	1.3292	1.4081	1.6083	1.9768
Kurtosis of y	5.4000	5.6181	6.5418	8.4693
Situation I				
$T_{SA_i}^c,\ i = 1,7$	106.0998	105.5777	105.4108	104.7131
$\mathbf{T_{SA_4}^c}$	**106.1067**	**105.5846**	**105.4176**	**104.7198**
$T_{R_1}^c$	100.2267	99.3709	98.3220	95.5744
$T_{DP_1}^c/T_{S_i}^c,\ i = 1,4$	105.983	105.46	105.3005	104.6141
Situation II				
$T_{SA_i}^c,\ i = 2,8$	115.0805	113.7576	113.4881	111.7679
$\mathbf{T_{SA_5}^c}$	**115.1003**	**113.7771**	**113.5075**	**111.7865**
$T_{R_1}^c$	97.5580	95.7573	93.4999	87.5916
$T_{DP_2}^c/T_{S_i}^c,\ i = 2,5$	114.9637	113.6399	113.3778	111.6689
Situation III				
$T_{SA_i}^c,\ i = 3,9$	108.1050	107.4685	107.3868	106.5397
$\mathbf{T_{SA_6}^c}$	**108.1161**	**107.4795**	**107.3977**	**106.5502**
$T_{R_1}^c$	97.3431	96.3414	95.0160	91.2945
$T_{DP_3}^c/T_{S_i}^c,\ i = 3,6$	107.9882	107.3508	107.2765	106.4406
$x^* \sim Exp(3.0)$				
$y^* \sim Exp(2.0)$				
Skewness of y	1.4612	1.3814	1.3734	1.4769
Kurtosis of y	5.9268	5.4885	5.4119	5.8395
Situation I				
$T_{SA_i}^c,\ i = 1,7$	106.2035	106.5419	106.7350	105.8748
$\mathbf{T_{SA_4}^c}$	**106.2096**	**106.5481**	**106.7411**	**105.8808**
$T_{R_1}^c$	99.7239	100.2227	100.2960	98.5947
$T_{DP_1}^c/T_{S_i}^c,\ i = 1,4$	106.1075	106.4467	106.6419	105.7829
Situation II				
$T_{SA_i}^c,\ i = 2,8$	112.9619	112.5542	112.8376	111.6621
$\mathbf{T_{SA_5}^c}$	**112.9805**	**112.5726**	**112.856**	**111.6799**
$T_{R_2}^c$	89.3047	88.4097	88.1260	85.9139
$T_{DP_2}^c/T_{S_i}^c,\ i = 2,5$	112.8659	112.4590	112.7445	111.5702
Situation III				
$T_{SA_i}^c,\ i = 3,9$	105.3578	105.7988	105.8817	105.2694
$\mathbf{T_{SA_6}^c}$	**105.3684**	**105.8092**	**105.8920**	**105.2796**
$T_{R_3}^c$	87.5260	88.2363	87.8974	86.9790
$T_{DP_3}^c/T_{S_i}^c,\ i = 3,6$	105.2617	105.7036	105.7886	105.1775

$j = 4, 5, 6$ are best among the proposed classes of imputation methods under situations I, II and III.

Table 2: *PRE* of proposed combined estimators at P=0.6

ρ_{xy}	0.6	0.7	0.8	0.9
T_m^c	100	100	100	100
$x^* \sim \Gamma(0.5, 1.5)$				
$y^* \sim \Gamma(1.5, 2)$				
Skewness of y	1.3292	1.4081	1.6083	1.9768
Kurtosis of y	5.4000	5.6181	6.5418	8.4693
Situation I				
$T_{SA_i}^c,\ i = 1, 7$	109.3365	108.5032	108.2434	107.1516
$\mathbf{T_{SA_4}^c}$	**109.3437**	**108.5104**	**108.2505**	**107.1584**
$T_{R_1}^c$	100.3407	99.0588	97.5032	93.5045
$T_{DP_1}^c/T_{S_i}^c,\ i = 1, 4$	109.2587	108.4248	108.1698	107.0855
Situation II				
$T_{SA_i}^c,\ i = 2, 8$	115.3909	114.0088	113.6905	111.9040
$\mathbf{T_{SA_5}^c}$	**115.4041**	**114.0218**	**113.7034**	**111.9163**
$T_{R_2}^c$	98.2571	96.3813	94.0605	88.0659
$T_{DP_2}^c/T_{S_i}^c,\ i = 2, 5$	115.3131	113.9304	113.6170	111.8379
Situation III				
$T_{SA_i}^c,\ i = 3, 9$	105.1641	104.7748	104.7360	104.2151
$\mathbf{T_{SA_6}^c}$	**105.1690**	**104.7796**	**104.7407**	**104.2197**
$T_{R_3}^c$	97.9303	97.2720	96.3820	93.8046
$T_{DP_3}^c/T_{S_i}^c,\ i = 3, 6$	105.0863	104.6964	104.6625	104.1491
$x^* \sim Exp(3.0)$				
$y^* \sim Exp(2.0)$				
Skewness of y	1.4612	1.3814	1.3734	1.4769
Kurtosis of y	5.9268	5.4885	5.4119	5.8395
Situation I				
$T_{SA_i}^c,\ i = 1, 7$	109.5194	110.0755	110.3901	109.0041
$\mathbf{T_{SA_4}^c}$	**109.5258**	**110.0819**	**110.3965**	**109.0104**
$T_{R_1}^c$	99.5862	100.3351	100.4455	97.9046
$T_{DP_1}^c/T_{S_i}^c,\ i = 1, 4$	109.4554	110.0121	110.3281	108.9429
Situation II				
$T_{SA_i}^c,\ i = 2, 8$	113.7547	113.6404	113.9884	112.5204
$\mathbf{T_{SA_5}^c}$	**113.7669**	**113.6525**	**114.0006**	**112.5322**
$T_{R_2}^c$	91.4958	91.0133	90.8178	88.2189
$T_{DP_2}^c/T_{S_i}^c,\ i = 2, 5$	113.6907	113.577	113.9264	112.4592
Situation III				
$T_{SA_3}^c$	103.4189	103.5222	103.5633	103.3902
$\mathbf{T_{SA_6}^c}$	**103.4436**	**103.5267**	**103.5677**	**103.4246**
$T_{R_3}^c$	90.6449	90.7375	90.4534	89.9166
$T_{DP_3}^c/T_{S_i}^c,\ i = 3, 6$	103.3749	103.4588	103.5013	103.3590

3. The PRE of the proposed combined and separate classes of imputation methods $y_{.i_{SA_j}}^c$, $y_{.i_{SA_j}}^s$, $j = 1, 2, ..., 9$ and their conventional counterparts under situations I, II and III are contrary to the asymmetry which is similar

Table 3: PRE of proposed separate estimators at P=0.4

ρ_{xy}	0.6	0.7	0.8	0.9
T_m^s	100	100	100	100
$x^* \sim \Gamma(0.5, 1.5)$				
$y^* \sim \Gamma(1.5, 2)$				
Skewness of y	1.3292	1.4081	1.6083	1.9768
Kurtosis of y	5.4000	5.6181	6.5418	8.4693
Situation I				
$T_{SA_i}^s,\ i = 1,7$	106.0962	105.5742	105.4076	104.7104
$\mathbf{T_{SA_4}^s}$	**106.1027**	**105.5807**	**105.4141**	**104.7167**
$T_{R_1}^s$	100.6188	99.7795	98.7804	96.1227
$T_{DP_1}^s/T_{S_i}^s,\ i = 1,4$	105.983	105.4600	105.3005	104.6141
Situation II				
$T_{SA_i}^s,\ i = 2,8$	115.0769	113.7541	113.4850	111.7652
$\mathbf{T_{SA_5}^s}$	**115.0954**	**113.7724**	**113.5031**	**111.7826**
$T_{R_2}^s$	98.5741	96.7889	94.6180	88.8216
$T_{DP_2}^s/T_{S_i}^s,\ i = 2,5$	114.9637	113.6399	113.3778	111.6689
Situation III				
$T_{SA_i}^s,\ i = 3,9$	108.1014	107.4650	107.3836	106.5370
$\mathbf{T_{SA_6}^s}$	**108.1119**	**107.4753**	**107.3939**	**106.5469**
$T_{R_3}^s$	81.4466	80.3429	77.9883	72.6504
$T_{DP_3}^s/T_{S_i}^s,\ i = 3,6$	107.9882	107.3508	107.2765	106.4406
$x^* \sim Exp(3.0)$				
$y^* \sim Exp(2.0)$				
Skewness of y	1.4612	1.3814	1.3734	1.4769
Kurtosis of y	5.9268	5.4885	5.4119	5.8395
Situation I				
$T_{SA_i}^s,\ i = 1,7$	106.2010	106.5393	106.7324	105.8724
$\mathbf{T_{SA_4}^s}$	**106.2067**	**106.5451**	**106.7382**	**105.8781**
$T_{R_1}^s$	100.0305	100.5150	100.5927	98.9217
$T_{DP_1}^s/T_{S_i}^s,\ i = 1,4$	106.1075	106.4467	106.6419	105.7829
Situation II				
$T_{SA_i}^s,\ i = 2,8$	112.4594	112.5516	112.8350	111.6596
$\mathbf{T_{SA_5}^s}$	**112.4770**	**112.5691**	**112.8525**	**111.6766**
$T_{R_2}^s$	90.1295	89.1992	88.9243	86.7481
$T_{DP_2}^s/T_{S_i}^s,\ i = 2,5$	112.8659	112.4590	112.7445	111.5702
Situation III				
$T_{SA_i}^s,\ i = 3,9$	105.3552	105.7962	105.8791	105.2670
$\mathbf{T_{SA_6}^s}$	**105.3654**	**105.8061**	**105.889**	**105.2767**
$T_{R_3}^s$	74.6746	73.9110	73.4349	71.8010
$T_{DP_3}^s/T_{S_i}^s,\ i = 3,6$	105.2617	105.7036	105.7886	105.1775

to the results of McIntyre ([20]), Dell and Clutter ([12]) and Bhushan and Kumar ([3]) where they chose a wide range of skewed distributions and concluded that the asymmetry shows adverse effect over the efficiency of

Table 4: PRE of proposed separate estimators at P=0.6

ρ_{xy}	0.6	0.7	0.8	0.9
T_m^s	100	100	100	100
$x^* \sim \Gamma(0.5, 1.5)$				
$y^* \sim \Gamma(1.5, 2)$				
Skewness of y	1.3292	1.4081	1.6083	1.9768
Kurtosis of y	5.4000	5.6181	6.5418	8.4693
Situation I				
$T_{SA_i}^s,\ i=1,7$	109.3341	108.5009	108.2413	107.1498
$\mathbf{T_{SA_4}^s}$	**109.3409**	**108.5076**	**108.2479**	**107.1562**
$T_{R_1}^s$	100.9318	99.6695	98.1811	94.2940
$T_{DP_1}^s/T_{S_i}^s,\ i=1,4$	109.2587	108.4248	108.1698	107.0855
Situation II				
$T_{SA_i}^s,\ i=2,8$	115.3885	114.0065	113.6884	111.9022
$\mathbf{T_{SA_5}^s}$	**115.4008**	**114.0187**	**113.7005**	**111.9138**
$T_{R_2}^s$	99.2688	97.4083	95.1744	89.2927
$T_{DP_2}^s/T_{S_i}^s,\ i=2,5$	115.3131	113.9304	113.6170	111.8379
Situation III				
$T_{SA_i}^s,\ i=3,9$	105.1618	104.7725	104.7339	104.2133
$\mathbf{T_{SA_6}^s}$	**105.1663**	**104.7770**	**104.7383**	**104.2176**
$T_{R_3}^s$	75.3238	74.4041	72.0020	66.8509
$T_{DP_3}^s/T_{S_i}^s,\ i=3,6$	105.0863	104.6964	104.6625	104.1491
$x^* \sim Exp(3.0)$				
$y^* \sim Exp(2.0)$				
Skewness of y	1.4612	1.3814	1.3734	1.4769
Kurtosis of y	5.9268	5.4885	5.4119	5.8395
Situation I				
$T_{SA_i}^s,\ i=1,7$	109.5177	110.0737	110.3884	109.0025
$\mathbf{T_{SA_4}^s}$	**109.5238**	**110.0798**	**110.3944**	**109.0084**
$T_{R_1}^s$	100.0457	100.7759	100.8935	98.3895
$T_{DP_1}^s/T_{S_i}^s,\ i=1,4$	109.4554	110.0121	110.3281	108.9429
Situation II				
$T_{SA_i}^s,\ i=2,8$	113.7531	113.8386	113.9867	112.5188
$\mathbf{T_{SA_5}^s}$	**113.7646**	**113.8502**	**113.9983**	**112.5300**
$T_{R_2}^s$	92.3101	91.7963	91.6107	89.0465
$T_{DP_2}^s/T_{S_i}^s,\ i=2,5$	113.6907	113.7770	113.9264	112.4592
Situation III				
$T_{SA_i}^s,\ i=3,9$	103.4372	103.5205	103.5615	103.4185
$\mathbf{T_{SA_6}^s}$	**103.4417**	**103.5247**	**103.5658**	**103.4228**
$T_{R_3}^s$	70.0312	69.5489	69.0562	67.4148
$T_{DP_3}^s/T_{S_i}^s,\ i=3,6$	103.3749	103.4588	103.5013	103.3590

the estimators.

4. The suggested combined classes of imputation methods $y^c_{.i_{SA_j}}$, $j = 1, 2, ..., 9$

are superior than the suggested separate classes of imputation methods $y^s_{.i_{SA_j}}$, $j = 1, 2, ..., 9$ in situations I, II and III.

Therefore, due to the dominance of the proposed imputation methods over the existing imputation methods, we recommend them to survey persons for their real life problems.

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Appendixes A-B-C. Supplementary file

Supplementary data to this article can be found online.